

A Bayesian Approach to Selection Effects Modelling

Wahidur Rahman

Co-Authors: Roberto Trotta, David van Dyk, Vincent Chen, Evan Tey and Timothy Lucas Makinen

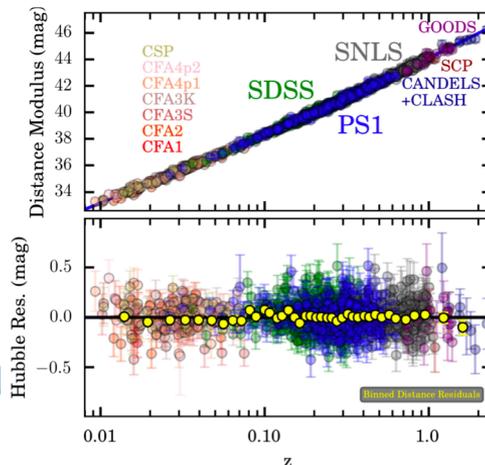
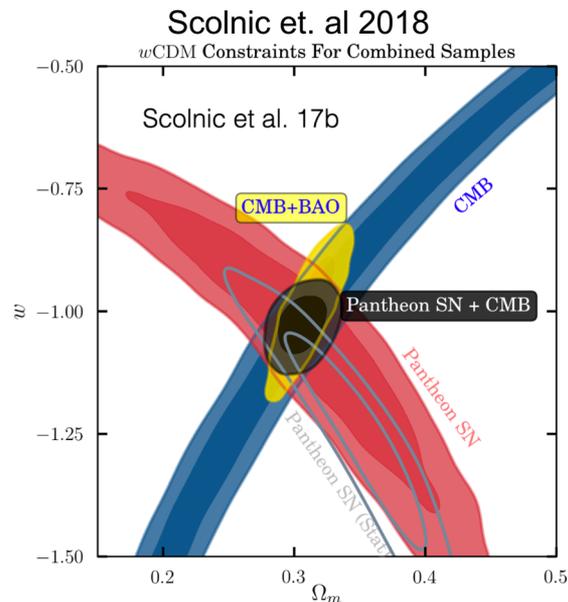
Overview

- Intro and Analysis of SN1a within the Bayesian Hierarchical Framework
BAHAMAS
 - Current problems with modelling Selection Effects of SN1a
 - Modelling of Selection Effects with Bayesian Methods
-

Supernova Cosmology

In the era of precision cosmology we are able to place tight constraints on our cosmology which SNe contribute to.

Currently limited by systematics of SNe data.



Standard Procedure

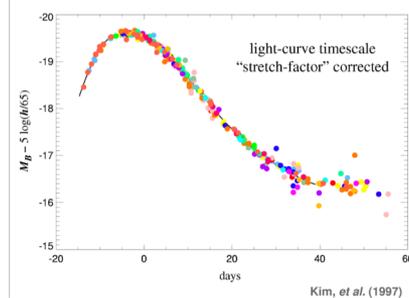
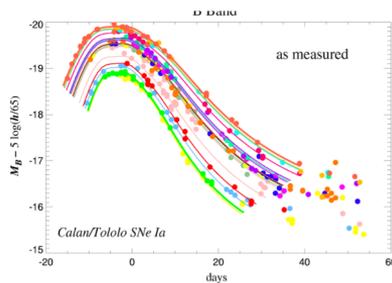
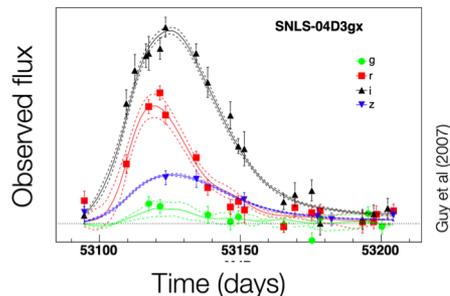
Spectroscopic Confirmation



Standardization Procedure



Chi-squared fit and minimization

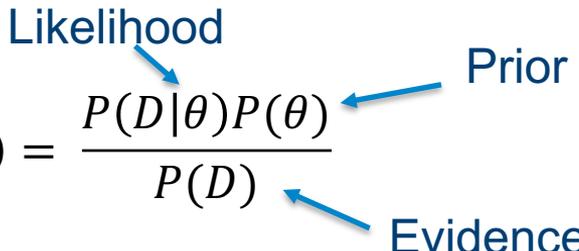


$$\chi^2 = \sum_i \frac{(\mu(z_i, C) - [m_{B,obs,i} - M_0 + \alpha x_{1,obs,i} + \beta c_{obs,i}])^2}{\sigma_{int}^2 + \sigma_{\mu_i}^2 + \sigma_{fit}^2}$$

Alternative to Chi-Squared Fit - BAHAMAS

BAyesian HierARchical Modelling for the Analysis of Supernova Cosmology

- Fits into the last step and based on Bayesian Statistics and designed to allow easy to access features such as model comparison and hierarchical modelling.
- Other Bayesian models have appeared for this step which include *UNITY* by Rubin et al (2015) and *Simple-BayesN* by Mandel et al (2017) and *Steve* by Hinton et al (2019).

$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)}$$


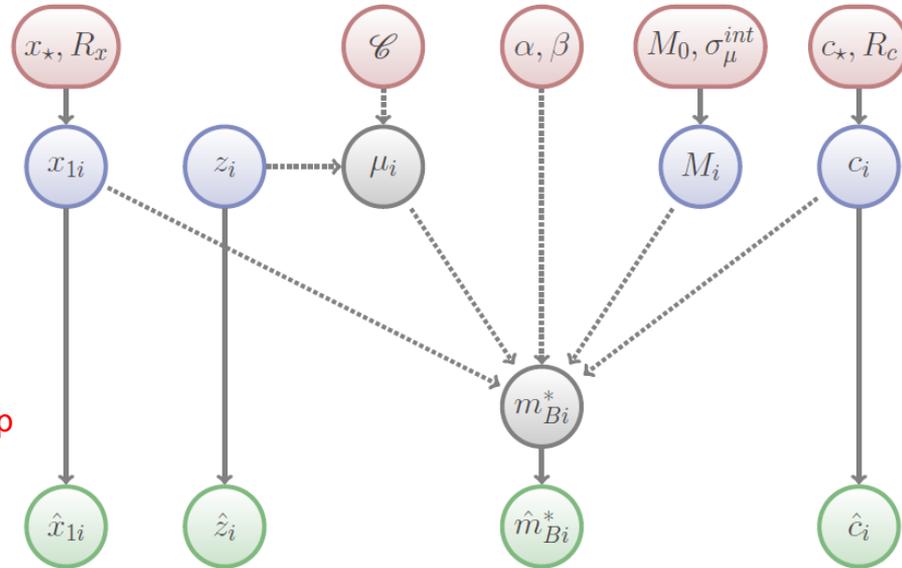
The diagram illustrates the components of the Bayesian equation. Three blue arrows point from text labels to parts of the equation: 'Likelihood' points to the numerator term $P(D|\theta)$, 'Prior' points to the numerator term $P(\theta)$, and 'Evidence' points to the denominator term $P(D)$.

BAHAMAS

Using a Hierarchical Modelling, we can relate the observed parameters to the parameters of interest and the true latent values which can be integrated out later in the likelihood.

$$\begin{aligned} p(\text{data}|\text{params}) &\propto \int p(\text{data}, \text{true}, \text{pop} | \text{params}) d\text{true} d\text{pop} \\ &= \int p(\text{data}|\text{true}) p(\text{true}|\text{pop}) p(\text{pop}) d\text{true} d\text{pop} \end{aligned}$$

March et al (2011)



Application of BAHAMAS Real Data - JLA

Set $w = -1$

JLA Alone

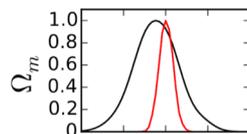
$$\Omega_M = 0.340 \pm 0.101$$

$$\Omega_K = 0.119 \pm 0.249$$

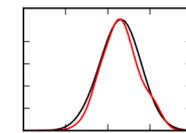
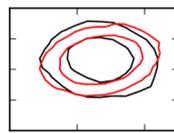
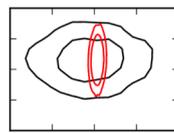
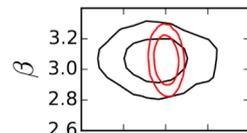
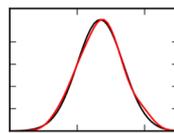
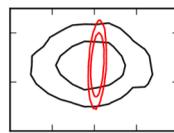
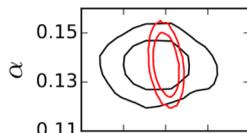
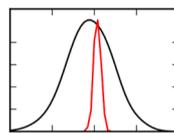
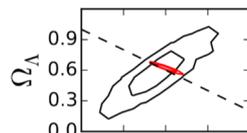
+Planck

$$\Omega_M = 0.399 \pm 0.027$$

$$\Omega_K = -0.024 \pm 0.008$$



Shariff et al (2015)



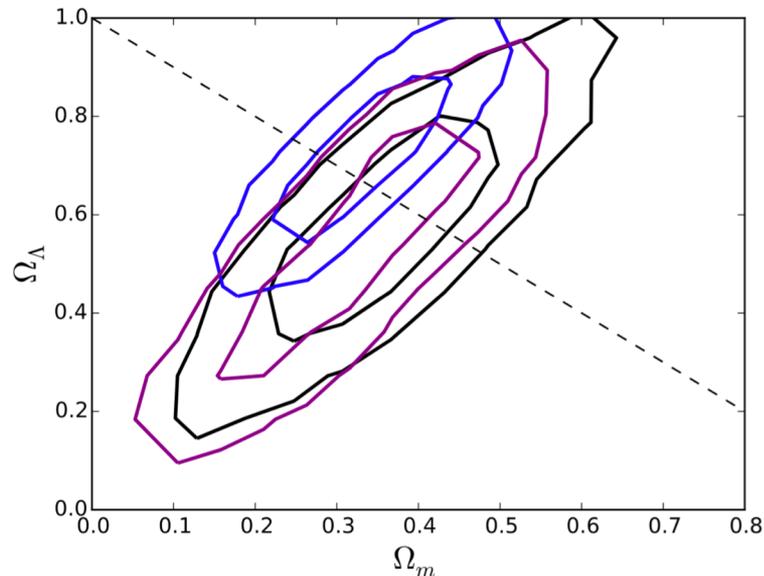
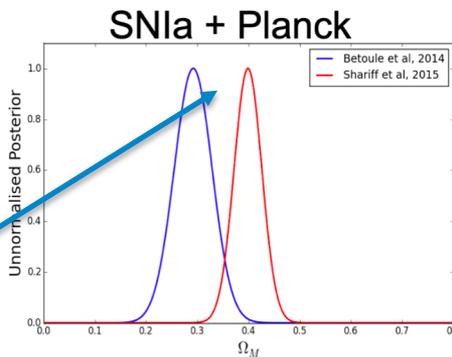
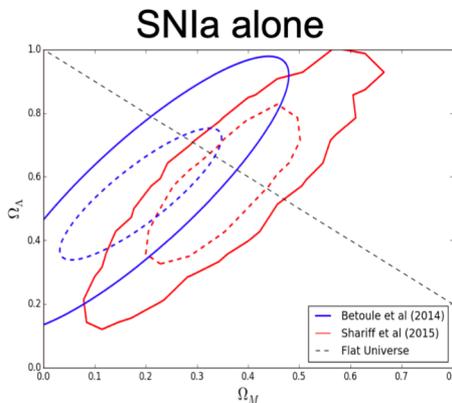
JLA (Betoule et al, 2014)
JLA+PLANCK 2015

Imperial College
London

Issues

Sensitive to incorrect
model specification

2.8 sigma tension



Blue: Statistical

Purple: Stat + m_B syst

Black: Combined

Selection Effects – The Problem

Problem: There is a Malmquist bias present in SNIa, which is both magnitude and colour dependent, as intrinsically fainter and redder SNIa are harder to detect.

Current Method: solution is to simulate data and apply adhoc corrections to the original data such as the BBC method (Kessler and Scolnic, 2017).

Downsides: Dependent on the cosmological parameters used for simulation.

Uncertainty in correction is difficult to incorporate.

Solution – A Bayesian Way

SNIs are indicated as selected or missed using an indicator function on some data $Y = \{Y_1, \dots, Y_n\}$ of which N objects are observed with $N < n$

$$I = \begin{cases} 1 & \text{if SNe } Y_i \text{ is observed} \\ 0 & \text{if SNe } Y_i \text{ is unobserved} \end{cases} \quad \begin{aligned} y_{obs} &= P(Y_i | I = 1, i \in [1, \dots, n]), \\ y_{miss} &= P(Y_i | I = 0, i \in [1, \dots, n]) \end{aligned}$$

$$p(y_{obs}, I | \theta, \phi) = \int p(Y, I | \theta, \phi) dy_{miss}, \quad \theta = \text{cosmo params}, \quad \phi = \text{selection params}$$

We can construct our posterior conditional on the observed data using Bayes Theorem and combine with our original likelihood. Many ways to model the Selection Function $P(Y_i, I = 1 | \theta, \phi)$. We use a Normal CDF which is linear in the cut parameters with smoothly varying selection efficiency. Cut parameters often inferred from simulations from tools such as SNANA.

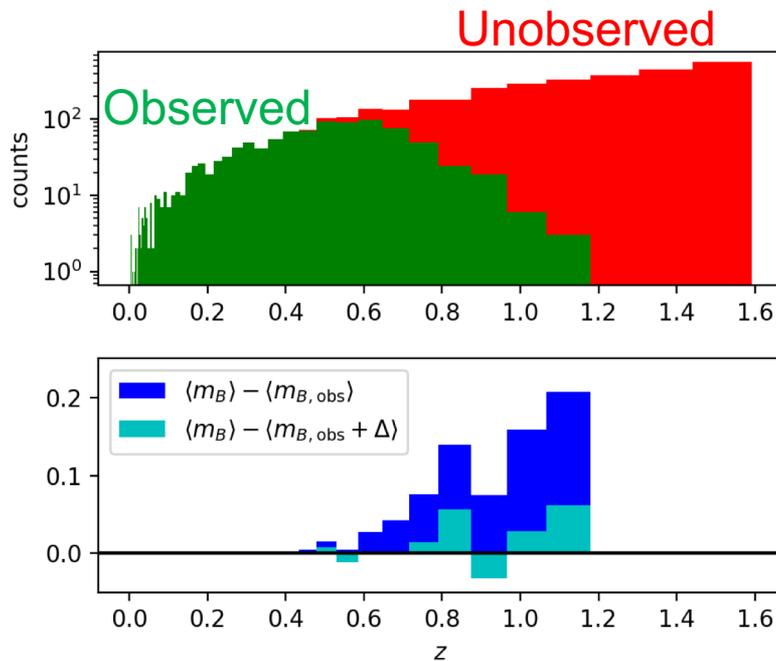
Solution – A Bayesian Way

Likelihood is similar to that used by Rubin et al (2015), but differs due to some fundamental assumptions about how we model the redshift distribution of unobserved SNe.

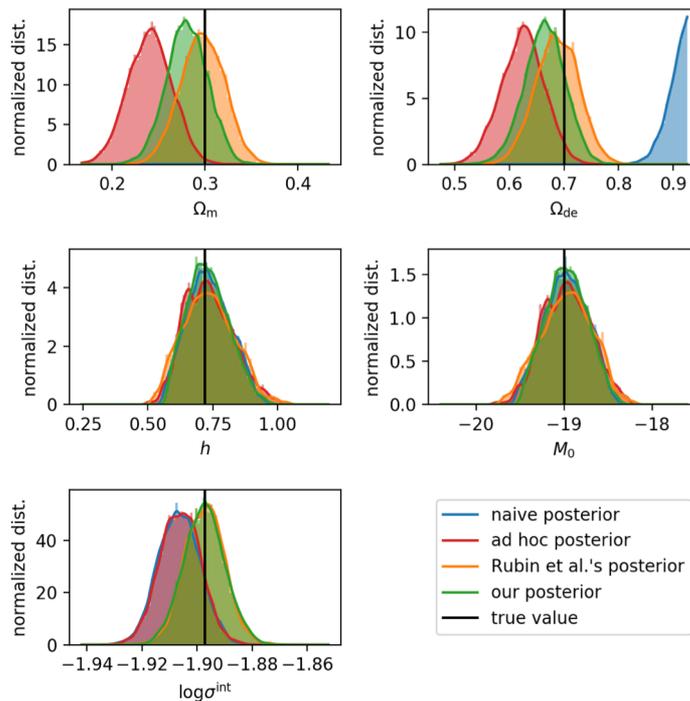
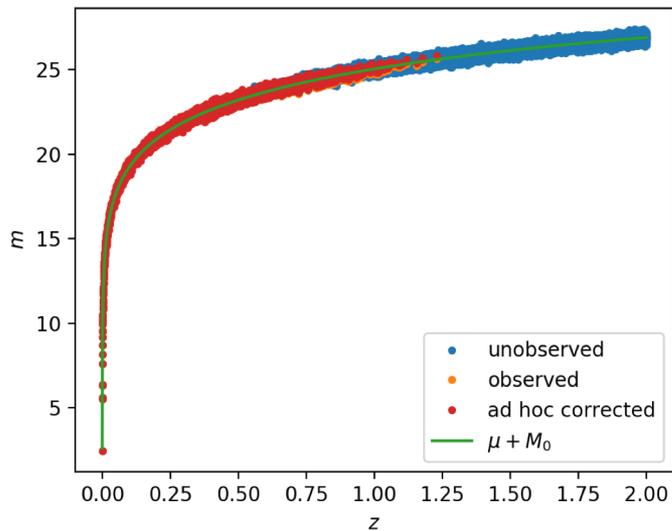
$$p(n_k) \propto 1/n_k$$

Prior is on the **total** unobserved supernova and marginalize out the redshift. R15 model implicitly assumes the same prior on observed and unobserved SNe. Leads to a different Likelihood, but it can be shown that it is related to our likelihood via:

$$P(y_{obs}|\theta, \phi) = \int L_{r15} P(z|\theta, \phi) dz$$

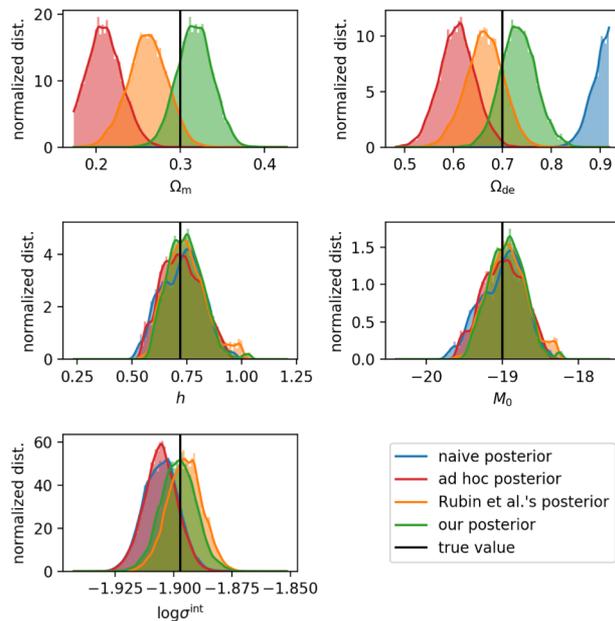


Solution – when the selection function is known and correct



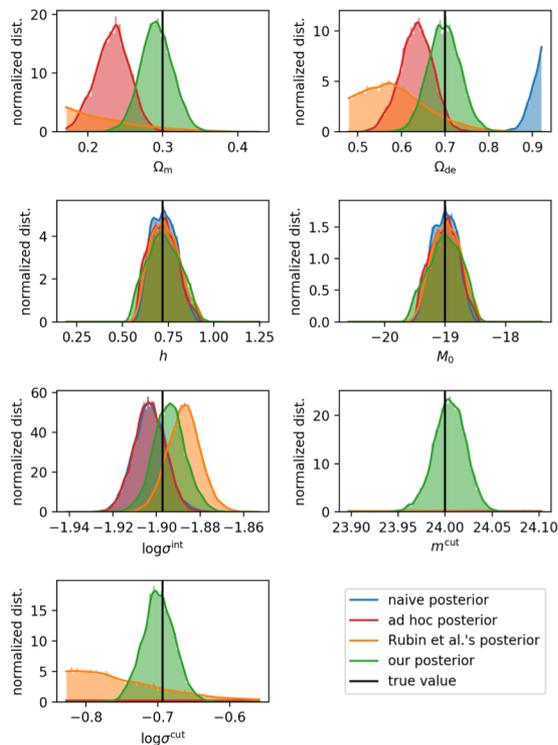
Solution – when the selection function is known and incorrect

Cut on m_B biased low $\sim 0.1\text{Mag}$



Solution – when the selection function is unknown

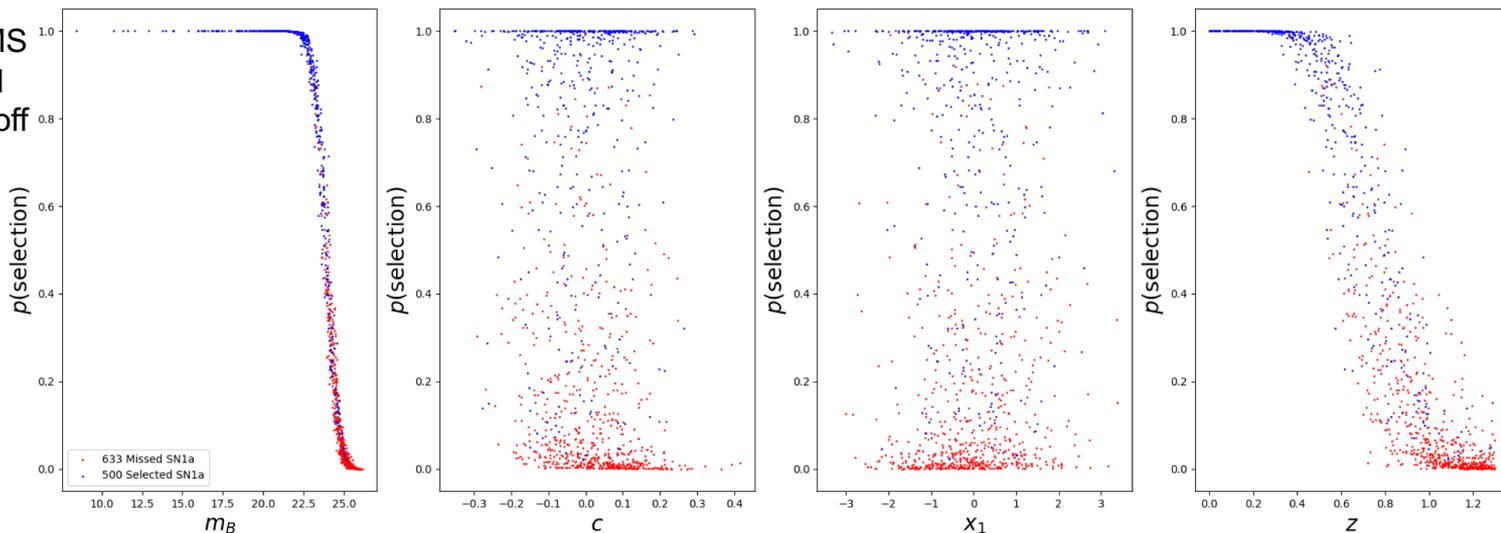
Perform inference
over the selection
parameters to solve
this issue



Solution – A Realistic Case

Simulated SNLS-like
selection efficiency
based off SNANA-SIMS
and applied to forward
modelled data based off
BAHAMAS

Simulated Selection Efficiency Based on SNLS-like SNANA simulations.

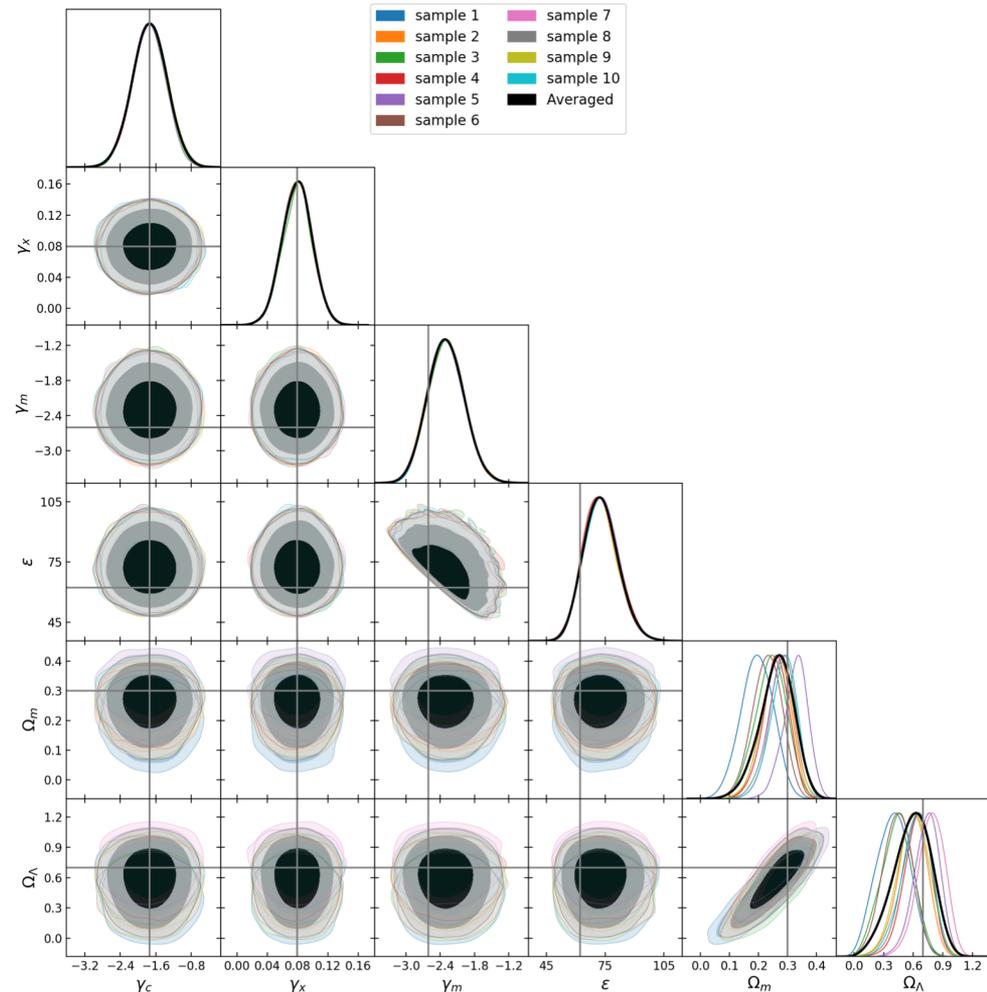


Ongoing Work

Solution – A Realistic Case

Using SNLS-like cuts on forward-generated data from the BAHAMAS model, we can assess the method on more realistic SNIa data

Ongoing Work



Conclusion:

Our method performs demonstrably well in the (realistic) scenario where the selection function is unknown.

BAHAMAS model still incomplete – work needed to incorporate other sources of systematic effects in model such host-metallicity and star formation rates.

Thank you
